## **CHAPTER 9**

## A note on computing with SPSS (Version 11.5):

We use the SPSS software to fit the nonlinear regression models of Chapter 9. SPSS works like a spreadsheet program. We enter the data into the various columns of the spreadsheet and use the tabs: Analyze > Regression > Nonlinear. We write out the model equation and specify initial parameter values. We can save the fitted values and the residuals (also the derivatives of the objective function) into columns of the worksheet.

Several options for the iterative nonlinear estimation procedure are available. In the following examples we have used the Levenberg-Marquardt algorithm. Options for specifying the number of iterations and various convergence cutoffs are available. See the SPSS on-line help for further discussion and examples.

9.1 A graph of the leaf area against the age of the palm tree is given below.



Note that there is not an abundance of data points to determine the model. The graph indicates that the relationship between leaf area and age is not linear; a quadratic component needs to be added to the model. The estimation results for the quadratic model  $y = \beta_0 + \beta_1 Age + \beta_2 Age^2 + \varepsilon$  (Minitab output) is shown below. The quadratic coefficient is clearly needed; the estimate of the coefficient for Age\*\*2 is -0.09616, with a significant t-ratio of -4.95.

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Regression Analysis: Area (square meters) versus Age, Age\*\*2The regression equation is<br/>Area (square meters) = -0.123 + 2.15 Age -0.0962 Age\*\*2PredictorCoefSE CoefTPredictorCoefSE CoefTPredictor0.12340.7334-0.17Age2.14960.25948.29Age\*\*2-0.096160.01942-4.95S = 0.7096R-Sq = 96.6%R-Sq(adj) = 95.8%Analysis of VarianceSMSFSourceDFSSMSFRegression2128.07164.036127.19Residual Error94.5310.503Total11132.60310

Rasch/Sedlacek use the Gompertz model  $y = \mu + \varepsilon = \alpha \exp[-\beta \exp(-\gamma Age)] + \varepsilon$  with parameters  $\alpha > 0, \beta > 0, \gamma > 0$ . Before fitting this model, we need to determine suitable starting values for the iterative nonlinear parameter estimation. The graph indicates that the saturation level for large values of Age is about 15. Hence a suitable starting value for  $\alpha$  is given by 15. For Age = 1, the response is about 2; for Age = 5, the response is roughly 7. The model equation implies  $-\beta \exp(-\gamma) = \ln(2/15)$  and  $-\beta \exp(-5\gamma) = \ln(7/15)$ . This implies  $\exp(4\gamma) = [\ln(2/15)]/[\ln(7/15)]$  and  $\gamma = \{\ln[\ln(2/15))/\ln(7/15)]\}/4 \approx 0.25$ . Finally,  $-\beta \exp(-\gamma) = \ln(2/15)$  and  $\beta = -\ln(2/15)\exp(\gamma) \approx 2.6$ . The starting values  $\alpha = 15$ ,  $\beta = 2.6$  and  $\gamma = 0.25$  are used in the SPSS nonlinear regression routine. The (SPSS) outcome is given below:

Iteration	Residual SS	A	В	С
-	10 5000000	1 - 0000000		05000000
1	12.59000092	15.0000000	2.60000000	.250000000
1.1	15.64377972	11.3812687	2.34691685	.336739045
1.2	6.515778841	13.4641276	2.14122037	.271436482
2	6.515778841	13.4641276	2.14122037	.271436482
2.1	6.243186484	12.0109653	2.42204992	.359733910
3	6.243186484	12.0109653	2.42204992	.359733910
3.1	5.136619171	12.4921144	2.50012161	.359000316
4	5.136619171	12.4921144	2.50012161	.359000316
4.1	5.136518308	12.4937910	2.49764737	.358922047
5	5.136518308	12.4937910	2.49764737	.358922047
5.1	5.136518286	12.4936881	2.49773050	.358935226

Run stopped after 11 model evaluations and 5 derivative evaluations. Iterations have been stopped because the relative reduction between successive residual sums of squares is at most SSCON = 1.000E-08

Nonlinear Reg	gression Sum	mary	Statis	tics	Depend	ent Va	riable A
Source		DF	Sum of	Squar	es Mean	Square	
Regression Residual Uncorrected (Corrected	d Total Total)	3 9 12 11	102 102 13	3.4341 5.1365 8.5707 2.6026	8 34 2 0 9	1.1447 .5707	3 2
R squared :	= 1 - Residu	al SS	S / Cor	rected	SS =	.9612	6
Parameter	Estimate	Asy Sto	mptoti l. Erro	c r	Asympto Confidenc Lower	tic 95 e Inte U	% rval pper
A ( $lpha$ )	12.49368805	7.	683789	772 10	.94684812	7 14.0	40527986
в ( $eta$ )	2.49773049	7	.440644	1 1	.50092433	8 3.4	94536656
С (ү)	.35893522	6.	066769	083	.20789306	7.5	09977385
Asymptotic	Correlation	Matr	ix of	the Pa	rameter E	stimat	es
	A		В	С			
A B C	1.0000 4983 8306	498 1.000 .833	33 - 10 39 1	.8306 .8339 .0000			

Dependent Variable AREA

The estimate of  $\alpha$  is 12.5; the estimate of  $\beta$  is 2.5, and the estimate of  $\gamma$  is 0.36. All estimates are statistically significant. There is a fair amount of correlation, especially between the estimates of  $\gamma$  and  $\alpha$  (-0.83) and the estimates of  $\gamma$  and  $\beta$  (0.83). The coefficient of determination (0.961) is similar to the  $R^2$  from the quadratic regression. There is little difference between the fits of the quadratic regression (which is linear in the parameters) and the Gompertz model (which is nonlinear in the parameters). Both models lead to similar fitted curves. One difference is that the fitted values for the Gompertz model increase with age to an asymptotic value, whereas the quadratic curve starts to decrease with age after having reached a maximum. However, over the observed age range the two fitted models are virtually indistinguishable.

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**9.2** A scatter plot of nitrate utilization versus light intensity is shown below. We use solid circles for day 1 observations, and triangles for day 2 observations. Furthermore, we have added some jitter to the light intensity in order to emphasize the differences between the measurements of day 1 and day 2. The day 2 measurements are slightly lower, especially at increasing light intensity.

Exercise 9.2: Plot of nitrate utilization against light intensity



<u>Michaelis-Menton model</u>: Nitrate utilization reaches an asymptote of about 20,000 for large light intensity. Letting x go to infinity in the model equation

$$\frac{\beta_1 x}{\beta_2 + x} = \frac{\beta_1}{1 + (\beta_2 / x)} \approx 20,000$$

leads to the starting value  $\beta_1 \approx 20,000$ . Furthermore, the average nitrate utilization at light intensity 2.2 is 1075. Solving the model equation with  $\beta_1 = 20,000$  leads to the starting value  $\beta_2 = 38.7$ .

Using these starting values in the SPSS nonlinear regression routine results in the following estimation results:

Nonlinear Regression Summary Statistics Dependent Variable NITRATE Source DF Sum of Squares Mean Square Regression26467226758.313233613379.15Residual4696536195.69322098612.94985Uncorrected Total486563762954.00(Corrected Total)472076766799.92 R squared = 1 - Residual SS / Corrected SS = .95352 Asymptotic 95 % Asymptotic Confidence Interval Std. Error Lower Upper Parameter Estimate в1 23582.527043 889.35646658 21792.345325 25372.708760 в2 34.243774004 3.427314571 27.344947587 41.142600421 Asymptotic Correlation Matrix of the Parameter Estimates B1 B2 1.0000 .8785 B1 в2 .8785 1.0000

Exponential rise model: Nitrate utilization reaches an asymptote of about 20,000 for large light intensity. Letting x go to infinity in the equation for the exponential rise model leads to the starting value  $\beta_1 \approx 20,000$ . The average nitrate utilization at light intensity 2.2 is 1075. Solving the model equation with  $\beta_1 = 20,000$  leads to the

starting value  $\beta_2 = -\frac{1}{2.2} \ln \left[ 1 - \frac{1075}{20,000} \right] = 0.025$ . Using these starting values in the

SPSS nonlinear regression program results in the estimation results:

Nonlinear Regression	Summary	Statistics	Dependent Variable NITRATE
Source	DF	Sum of Squares	Mean Square
Regression	2	6504309173.87	3252154586.93
Residual	46	59453780.1310	1292473.48111
Uncorrected Total	48	6563762954.00	
(Corrected Total)	47	2076766799.92	
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R squared = 1 - Residual SS / Corrected SS = .97137 Asymptotic 95 % Asymptotic Confidence Interval Std. Error Lower Upper Parameter Estimate 19014.305975 398.04663684 18213.079652 19815.532299 В1 в2 .030021624 .001629334 .026741945 .033301303 Asymptotic Correlation Matrix of the Parameter Estimates в2 В1 в1 1.0000 -.7393 -.7393 1.0000 В2

<u>Quadratic Michaelis-Menton model</u>: Starting with  $\beta_1 = 20,000$  and  $\beta_2 = 38.7$  (from the earlier Michaelis-Menton model) and a small value for the parameter in the quadratic component ( $\beta_3 = 0.1$ ) leads to the following results:

Nonlinear Re	gression Summ	nary Stat	istics	Depend	ent Variable	NITRATE
Source		DF Sum	of Squa	res Mean	Square	
Regression Residual Uncorrecte (Corrected	d Total d Total)	3 6520 45 4322 48 6563 47 2076	)540397. 22556.66 3762954. 5766799.	33 217351 54 96050 00 92	3465.78 1.25923	
R squared	= 1 - Residua	al SS / C	Correcte	d SS =	.97919	
Parameter	Estimate	Asympto Std. Er	otic cror	Asympto Confidenc Lower	tic 95 % e Interval Upper	
B1 B2 B3	66769.034924 137.82679758 .011281055	17585.50 43.73571 .00449	)4714 31 L2594 49 ∂6402	350.010284 .738550634 .002224837	102188.0595 225.9150445 .02033727	6 3 4
Asymptotic	Correlation M	Matrix of	the Pa	rameter Es	timates	
	B1	в2	в3			

B1	1.0000	.9964	.9941
В2	.9964	1.0000	.9856
В3	.9941	.9856	1.0000

<u>Modified exponential rise model</u>: Using  $\beta_1 = 20,000$  and  $\beta_2 = 0.025$  from the earlier exponential rise model and a small value for  $\beta_3 = 0.01$  leads to the following results:

Ν	onlinear F	Regression	Summary	Statis	tics	Dependent	Variable	NITRATE
	Source		DF	Sum of	Squares	Mean Squa	are	
	Regressic Residual Uncorrect (Correcte	on ced Total ed Total)	3 45 48 47	651911 446458 656376 207676	7089.28 64.7154 2954.00 6799.92	217303902 992130.3	9.76 2701	
	R squared	d = 1 - Res	idual S	S / Cor	rected S	s = .9	7850	
	Parameter	r Estimat	As: ce Sto	ymptoti d. Erro:	c Con r Lov	Asymptotic nfidence I wer	95 % nterval Upper	
	B1 B2 B3	33551.454 .018534 .003221	219 950 079 . 159 .	2.16877 0035721 0013385	11 14413 51 .013 59 .00	.103896 520 1339397 0525162	689.80454 .02572876 .00591715	3 1 5
	Asymptoti	lc Correlat	ion Mat	rix of	the Para	meter Estin	mates	
		B1	1	В2	в3			
	B1 B2 B3	1.0000 9898 .9948	98 1.00 97	98 00 - 41 1	.9948 .9741 .0000			

All four models lead to large  $R^2$ . The Michaelis-Menton and its quadratic extension lead to  $R^2$  of 0.954 and 0.979, respectively. Carrying out an F-test for the significance of the quadratic component in the Michaelis-Menton model leads to the F-statistic F = [96,536,195 - 43,222,556]/[43,222,556/45] = 55.5, which is highly significant. This shows that the quadratic extension represents a significant improvement.

Similarly, the exponential rise model and its extension lead to  $R^2$  of 0.971 and 0.979, respectively. The F-test for the significance of the extra component in the exponential rise model leads to the F-statistic

F = [59,453,780 - 44,645,864]/[44,645,864/45] = 14.9, which is also highly significant.

The extensions are beneficial. The modified Michaelis-Menton and the modified exponential rise models perform similarly. In the following graph we show the fit of the quadratic Michaelis-Menton model; the fitted values of the modified exponential rise model are virtually indistinguishable.



<u>Standard Michaelis-Menton model with an indicator for the change of day:</u> The final parameter estimates in the previous Michaelis-Menton model,  $\hat{\beta}_1 = 23,500$  and  $\hat{\beta}_2 = 34.2$ , are taken as the starting values in the iterative nonlinear estimation. Small values for the day indicator  $\alpha_1 = -1000, \alpha_2 = -1$  are used as the starting values for the two additional parameters. The estimation results are given below:

Nonlinear R	egression Summ	nary	Statisti	CS	Depende	ent V	Variable	NITRATE
Source		DF	Sum of Se	quares	Mean S	Squar	re	
Regression Residual Uncorrecto (Corrected	n ed Total d Total)	4 44 48 47	64772534 86509529 65637629 20767667	24.57 .4274 54.00 99.92	1619313 1966125	356. 5.668	14 381	
R squared	= 1 - Residua	al SS	5 / Corre	cted SS	5 =	.958	334	
		_		1	Asymptot	1C 9	15 8	
		Asy	mptotic	Cor	ntidence	e Int	erval	
Parameter	Estimate	Sto	d. Error	Lov	ver		Upper	
B1 B2 A1 A2	24743.334444 35.275400267 -2328.743446 -2.172827290	1241 4.6 1720 6.6	L.1211323 556586052 ).3472191 526226364	22242 25.890 -5795 -15.52	.019158 0667730 .875448 2710905	2724 44.6 1138 11.1	44.649730 560132803 3.3885565 81454466	) 3 7 5

Asymptotic Correlation Matrix of the Parameter Estimates

	B1	В2	Al	A2
В1	1.0000	.8810	7214	6191
B2	.8810	1.0000	6356	7028
A1	7214	6356	1.0000	.8781
A2	6191	7028	.8781	1.0000

The F-statistic for testing the null hypothesis  $\alpha_1 = \alpha_2 = 0$  is

F = [(96,536,195 – 86,509,529)/2]/[86,509,529/44] = 2.55. The probability value from the F(2,44) distribution is P[F(2,44)  $\ge$  2.55] = 1 – 0.91 = 0.09. Hence there is only weak evidence for including a day effect. The individual confidence intervals for  $\alpha_1$  and  $\alpha_2$  cover zero, which makes the individual interpretation of the two day-effect parameters difficult. These estimates are also quite correlated.

Quadratic Michaelis-Menton model with an indicator for the change of day: The final values from the earlier quadratic model  $\hat{\beta}_1 = 66,700, \hat{\beta}_2 = 138, \hat{\beta}_3 = 0.01$  and small values for the three parameters associated with the day indicators,  $\alpha_1 = -2000, \alpha_2 = -2, \alpha_3 = 0.001$ , are used as the starting values in the iterative nonlinear SPSS estimation. The estimation results are given below:

Run stopped after 10 model evaluations and 5 derivative evaluations. Iterations have been stopped because the relative reduction between successive residual sums of squares is at most SSCON = 1.000E-08

Nonlinear	Regression S	ummary	Statist	cics	Dependent	Variable	NITRATE
Source		DF	Sum of	Squares	Mean Squ	are	
Regressi Residual Uncorrec (Correct	on ted Total ed Total)	6 42 48 47	6531740 3202259 6563762 2076766	)362.05 91.9535 2954.00 5799.92	108862339 762442.6	3.67 6556	
R squared	. = 1 - Resid	ual SS	/ Corre	ected SS	= .98	458	
				i	Asymptotic	95 %	
		Asy	mptotic	c Coi	nfidence I	nterval	
Paramete	r Estimate	sto	d. Erroi	C Lov	wer	Upper	
B1 B2 A1 A2 B3	89797.9169 186.618624 -38897.786 -83.090781 .0162524	70 3754 45 89.9 90 3998 51 96.5 21 .0	40.34574 98455396 32.74885 72734645 00920728	$\begin{array}{rrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrrr$	.432096 16 2442558 36 86.2408 41 2944695 11 2328638	5557.4018 8.2148063 790.66703 2.1129065 .03483348	4 5 3 1
A.	.0004490			.020	0101001	· 01130200.	

Asymptotic	Correlation	Matrix	of	the	Parameter	Estimates
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	B1	В2	A1	A2	В3	A3
В1	1.0000	.9978	9389	9283	.9965	9252
в2	.9978	1.0000	9369	9303	.9913	9204
A1	9389	9369	1.0000	.9971	9356	.9953
A2	9283	9303	.9971	1.0000	9222	.9894
В3	.9965	.9913	9356	9222	1.0000	9285
A3	9252	9204	.9953	.9894	9285	1.0000

The F-statistic for testing the null hypothesis  $\alpha_1 = \alpha_2 = \alpha_3 = 0$  is given by F = [(43,222,556 - 32,022,591)/3]/[32,022,591/42] = 4.90. The probability value from the F(3,42) distribution is  $P[F(3,42) \ge 4.90] = 1 - 0.995 = 0.005$ , showing that the indicators for the day effect help explain the variation. Individually the three parameters are statistically insignificant and also highly correlated. This makes an individual interpretation of the estimates difficult.

The graph shown below compares the quadratic Michaelis-Menton model with and without the day indicator. The graph shows that the quadratic Michaelis-Menton model with a day indicator is capable of expressing the day differences.

Exercise 9.2: Quadratic Michaelis-Menton model with day indicator



## 9.3

<u>Model 1:</u> The logarithmic transformation of the first model leads to  $\ln(y) = \ln(\beta_0) + \beta_1 \ln(x_1) + \beta_2 \ln(x_2) + \ln(\varepsilon)$ 

A standard multiple linear regression of  $\ln(y)$  on  $\ln(x_1)$  and  $\ln(x_2)$  leads to the estimates of  $\alpha = \ln(\beta_0), \beta_1$ , and  $\beta_2$  The estimate of  $\beta_0$  can be obtained

from  $\beta_0 = \exp(\alpha)$ . When carrying out the regression with the transformed variables we need to assume that the error  $\ln(\varepsilon)$  satisfies the standard regression assumptions. <u>Model 2</u>: Taking the reciprocal of the response in the second model leads to

$$1/\mathbf{y} = \boldsymbol{\beta}_0 + \boldsymbol{\beta}_1 \mathbf{x} + \boldsymbol{\varepsilon}$$

A simple linear regression of (1/y) on  $x_1$  leads to the estimates of  $\beta_0, \beta_1$ .

<u>Model 3:</u> The reciprocal of the response and a subsequent logarithmic transformation leads to the model

 $\ln[(1/y) - 1] = \beta_0 + \beta_1 x_1 + \ln(\varepsilon)$ 

A simple linear regression of  $\ln[(1/y) - 1]$  on  $x_1$  leads to the estimates of  $\beta_0, \beta_1$ . We need to assume that the error  $\ln(\varepsilon)$  satisfies the standard regression assumptions.

**9.4** Search the literature.